

Streaming Algorithms for Optimal Generation of Random Bits

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Abstract—Generating random bits from a source of biased coins (the biased is unknown) is a classical question that was originally studied by von Neumann. There are a number of known algorithms that have asymptotically optimal information efficiency, namely, the expected number of generated random bits per input bit is asymptotically close to the entropy of the source. However, only the original von Neumann algorithm has a ‘streaming property’ - it operates on a single input bit at a time and it generates random bits when possible, alas, it does not have an optimal information efficiency.

The main contribution of this paper is an algorithm that generates random bit streams from biased coins, uses bounded space and runs in expected linear time. As the size of the allotted space increases, the algorithm approaches the information-theoretic upper bound on efficiency. In addition, we discuss how to extend this algorithm to generate random bit streams from m -sided dice or correlated sources such as Markov chains.

Index Terms—Random Number Generation, Biased Coins, Markov Chains, Streams.

I. INTRODUCTION

THE question of generating random bits from a source of biased coins dates back to von Neumann [8] who observed that when one focuses on a pair of coin tosses, the events HT and TH have the same probability (H is for ‘head’ and T is for ‘tail’) of being generated; hence, HT produces the output symbol 1 and TH produces the output symbol 0. The other two possible events, namely, HH and TT, are ignored, namely, they do not produce any output symbols. However, von Neumann’s algorithm is not optimal in terms of the number of random bits that are generated. This problem was solved, specifically, given a fixed number of biased coin tosses with unknown probability, it is well known how to generate random bits with asymptotically optimal efficiency, namely, the expected number of unbiased random bits generated per coin toss is asymptotically equal to the entropy of the biased coin [3]–[6]. However, these solutions, including Elias’s algorithm and Peres’s algorithm, can generate random bits only after receiving the complete input sequence (or a fixed number of input bits), and the number of random bits generated is a random variable.

We consider the setup of generating a “stream” of random bits; that is, whenever random bits are required, the algorithm reads new coin tosses and generates random bits dynamically. Our new streaming algorithm is more efficient (in the number

of input bits, memory and time) for producing the required number of random bits and is a better choice for implementation in practical systems. We notice that von Neumann scheme is the one which is able to generate a stream of random bits, but its efficiency is far from optimal. Our goal is to modify this scheme such that it can achieve the information-theoretic upper bound on efficiency. Specifically, we would like to construct a function $f : \{H, T\}^* \rightarrow \{0, 1\}^*$ which satisfies the following conditions:

- f generates a stream. For any two sequences of coin tosses $x, y \in \{H, T\}^*$, $f(x)$ is a prefix of $f(xy)$.
- f generates random bits. Let $X_k \in \{0, 1\}^*$ be the sequence of coin tosses inducing k bits; that is, $|f(X_k)| \geq k$ but for any strict prefix X of X_k , $|f(X)| \leq k$. Then the first k bits of $f(X_k)$ are independent and unbiased.
- f has asymptotically optimal efficiency. That is, for any $k > 0$,

$$\frac{E[|X_k|]}{k} \rightarrow \frac{1}{H(p)}$$

as $k \rightarrow \infty$, where $H(p)$ is the entropy of the biased coin [2].

We note that the von Neumann scheme uses only 3 states, i.e., a symbol in $\{\phi, H, T\}$, for storing state information. For example, the output bit is 1 if and only if the current state is H and the input symbol is T. In this case, the new state is ϕ . Similarly, the output bit is 0 if and only if the current state is T and the input symbol is H. In this case, the new state is ϕ . Our approach for generalizing von Neumann’s scheme is by increasing the memory (or state) of our algorithm such that we do not lose information that might be useful for generating future random bits. We represent the state information as a binary tree, called status tree, in which each node is labeled by a symbol in $\{\phi, H, T, 0, 1\}$. When a source symbol (a coin toss) is received, we modify the status tree based on certain simple rules and generate random bits in a dynamic way. This is the key idea in our algorithm; we call this approach the random-stream algorithm. In some sense, the random-stream algorithm is the streaming version of Peres’s algorithm. We show that this algorithm satisfies all three conditions above, namely, it can generate a stream of random bits with asymptotically optimal efficiency. In practice, we can reduce the space size by limiting the depth of the status tree. We will demonstrate that as the depth of the status tree increases, the efficiency of the algorithm quickly converges to the information-theoretic upper bound.

An extension of the question is to generate random bits or random-bit streams from an arbitrary Markov chain with

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unknown transition probabilities. This problem was first studied by Samuelson [7], and his algorithm was later improved by Blum [1]. Recently, we proposed the first known algorithm that runs in expected linear time and achieves the information-theoretic upper bound on efficiency [9]. In this paper, we briefly introduce the techniques of generating random-bit streams from Markov chains.

The rest of the paper is organized as follows. Section II presents our key result, the random-stream algorithm that generates random bit streams from arbitrary biased coins and achieves the information-theoretic upper bound on efficiency. In Section III, we generalize the random-stream algorithm to generate random bit streams from a source of a larger alphabet. An extension for Markov chains is provided in Section IV, followed by the concluding remarks.

II. THE RANDOM-STREAM ALGORITHM

A. Description

Many algorithms have been proposed for efficiently generating random bits from a fixed number of coins tosses, including Elias's algorithm and Peres's algorithm. However, in these algorithms, the input bits can be processed only after all of them have been received, and the number of random bits generated cannot be controlled. In this section, we focus on deriving a new algorithm, the *random-stream algorithm*, that generates a stream of random bits from an arbitrary biased-coin source and achieves the information-theoretic upper bound on efficiency. Given an application that requires random bits, the random-stream algorithm can generate random bits dynamically based on requests from the application.

While von Neumann's scheme can generate a stream of random bits from an arbitrary biased coin, its efficiency is far from being optimal. The main reason is that it uses minimal state information, recorded by a symbol of alphabet size three in $\{\phi, H, T\}$. The key idea in our algorithm is to create a binary tree for storing the state information, called a status tree. A node in the status tree stores a symbol in $\{\phi, H, T, 0, 1\}$. The following procedure shows how the status tree is created and is dynamically updated in response to arriving input bits. At the beginning, the tree has only a single root node labeled as ϕ . When reading a coin toss from the source, we modify the status tree based on certain rules. For each node in the status tree, if it receives a message (H or T), we do operations on the node. Meanwhile, this node may pass some new messages to its children. Iteratively, we can process the status tree until no more messages are generated. Specifically, let u be a node in the tree. Assume the label of u is $x \in \{\phi, H, T, 1, 0\}$ and it receives a symbol $y \in \{H, T\}$ from its parent node (or from the source if u is the root node). Depending on the values of x and y , we do the following operations on node u .

- 1) When $x = \phi$, set $x = y$.
- 2) When $x = 1$ or 0 , output x and set $x = y$.
- 3) When $x = H$ or T , we first check whether u has children. If it does not have, we create two children with label ϕ for it. Let u_l and u_r denote the two children of u .

- If $xy = HH$, we set $x = \phi$, then pass a symbol T to u_l and a symbol H to u_r .
- If $xy = TT$, we set $x = \phi$, then pass a symbol T to u_l and a symbol T to u_r .
- If $xy = HT$, we set $x = 1$, then pass a symbol H to u_l .
- If $xy = TH$, we set $x = 0$, then pass a symbol H to u_l .

We see that the node u passes a symbol $x + y \bmod 2$ to its left child and if $x = y$ it passes a symbol x to its right child.

Note that the timing is crucial that we output a node's label (when it is 1 or 0) only after it receives the next symbol from its parent or from the source. This is different from von Neumann's scheme where a 1 or a 0 is generated immediately without waiting for the next symbol. If we only consider the output of the root node in the status tree, then it is similar to von Neumann's scheme. And the other nodes correspond to the information discarded by von Neumann's scheme. In some sense, the random-stream algorithm can be treated as a "stream" version of Peres's algorithm. The following example is constructed for the purpose of demonstration.

Example 1. Assume we have a biased coin and our randomized application requires 2 random bits. Fig. 1 illustrates how the random-stream algorithm works when the incoming stream is *HTTTHT...* In this figure, we can see the changes of the status tree and the messages (symbols) passed throughout the tree for each step. We see that the output stream is 11...

Lemma 1. Let X be the current input sequence and let \mathcal{T} be the current status tree. Given \mathcal{T} and the bits generated by each node in \mathcal{T} , we can reconstruct X uniquely.

Proof: Let us prove this lemma by induction. If the maximum depth of the status tree is 0, it has only a single node. In this case, X is exactly the label on the single node. Hence the conclusion is trivial. Now we show that if the conclusion holds for all status trees with maximum depth at most k , then it also holds for all status trees with maximum depth $k + 1$.

Given a status tree \mathcal{T} with maximum depth $k + 1$, we let $Y \in \{0, 1\}^*$ denote the binary sequence generated by the root node, and $L, R \in \{H, T\}^*$ are the sequences of symbols received by its left child and right child. If the label of the root node is in $\{0, 1\}$, we add it to Y . According to the random-stream algorithm, it is easy to get that

$$|L| = |Y| + |R|.$$

Based on our assumption, L, R can be constructed from the left and right subtrees and the bits generated by each node in the subtree since their depths are at most k . We show that once L, R, Y satisfy the equality above, the input sequence X can be uniquely constructed from L, R, Y and α , where α is the label of the root node. The procedure is as follows: Let us start from an empty string for X and read symbols from L sequentially. If a symbol read from L is H , we read a bit from Y . If this bit is 1 we add HT to X , otherwise we add TH to X . If a symbol read from L is T , we read a symbol (H or T)

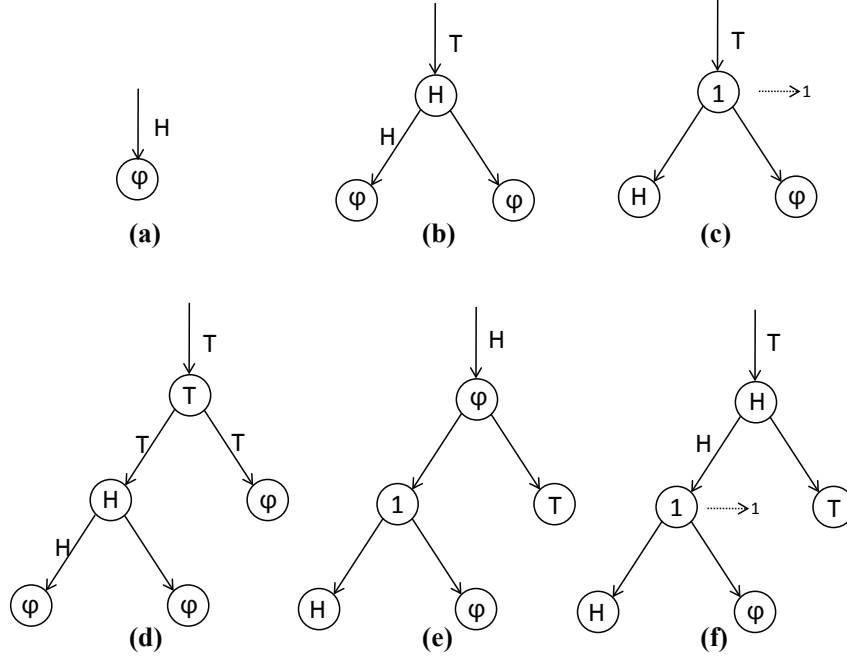


Fig. 1. An instance for generating 2 random bits using the random-stream algorithm.

from R . If this symbol is H we add HH to X , otherwise we add TT to X .

After reading all the elements in L, R and Y , the length of the resulting input sequence is $2|L|$. Now, we add α to the resulting input sequence if $\alpha \in \{H, T\}$. This leads to the final sequence X , which is unique. ■

Example 2. Let us consider the status tree in Fig. 1(f). And we know that the root node generates 1 and the first node in the second level generates 1. We can have the following conclusions iteratively.

- In the third level, the symbols received by the node with label H are H, and the node with label ϕ does not receive any symbols.
- In the second level, the symbols received by the node with label 1 are HTH, and the symbols received by the node with label T are T.
- For the root node, the symbols received are HTTHT, which accords with Example 1.

Let $f : \{H, T\}^* \rightarrow \{0, 1\}^*$ be the function of the random-stream algorithm. We show that this function satisfies all the three conditions described in the introduction. It is easy to see that the first condition holds, i.e., for any two sequences $x, y \in \{H, T\}^*$, $f(x)$ is a prefix of $f(xy)$, hence it generates streams. The following two theorems indicate that f also satisfies the other two conditions.

Theorem 2. Given a source of biased coin with unknown probability, the random-stream algorithm generates a stream of random bits, i.e., for any $k > 0$, if we stop running the algorithm after generating k bits then these k bits are independent and unbiased.

Let S_Y with $Y \in \{0, 1\}^k$ denote the set consisting of all

the binary sequences yielding Y . Here, we say that a binary sequence X yields Y if and only if $X[1 : |X| - 1]$ (the prefix of X with length $|X| - 1$) generates a sequence shorter than Y and X generates a sequence with Y as a prefix (including Y itself). To prove that the algorithm can generate random-bit streams, we show that for any distinct binary sequences $Y_1, Y_2 \in \{0, 1\}^k$, the elements in S_{Y_1} and those in S_{Y_2} are one-to-one mapping. The detailed proof is given in Subsection II-B.

Theorem 3. Given a biased coin with probability p being H, let n be the number of coin tosses required for generating k random bits in the random-stream algorithm, then

$$\lim_{k \rightarrow \infty} \frac{E[n]}{k} = \frac{1}{H(p)}.$$

The proof of Theorem 3 is based on the fact that the random-stream algorithm is as efficient as Peres's algorithm. The difference is that in Peres's algorithm the input length is fixed and the output length is variable. But in the random-stream algorithm the output length is fixed and the input length is variable. So the key of the proof is to connect these two cases. The detailed proof is given in Subsection II-C.

So far, we can conclude that the random-stream algorithm can generate a stream of random bits from an arbitrary biased coin with asymptotically optimal efficiency. However, the size of the binary tree increases as the number of input coin tosses increases. The longest path of the tree is the left-most path, in which each node passes one message to the next node when it receives two messages from its previous node. Hence, the maximum depth of the tree is $\log_2 n$ for n input bits. This linear increase in space is a practical challenge. Our observation is that we can control the size of the space by limiting the maximum depth of the tree – if a node's depth

reaches a certain threshold, it will stop creating new leaves. We can prove that this method correctly generates a stream of random bits from an arbitrary biased coin. We call this method the random-stream algorithm with maximum depth d .

Theorem 4. *Given a source of a biased coin with unknown probability, the random-stream algorithm with maximum depth d generates a stream of random bits, i.e., for any $k > 0$, if we stop running the algorithm after generating k bits then these k bits are independent and unbiased.*

The proof of Theorem 4 is a simple modification of the proof of Theorem 2, given in Subsection II-D. In order to save memory space, we need to reduce the efficiency. Fortunately, as the maximum depth increases, the efficiency of this method can quickly converge to the theoretical limit.

Example 3. *When the maximum depth of the tree is 0 (it has only the root node), then the algorithm is approximately von Neumann's scheme. The expected number of coin tosses required per random bit is*

$$\frac{1}{pq}$$

asymptotically, where $q = 1 - p$ and p is the probability for the biased coin being H.

Example 4. *When the maximum depth of the tree is 1, the expected number of coin tosses required per random bit is*

$$\frac{1}{pq + \frac{1}{2}(p^2 + q^2)(2pq) + \frac{1}{2}(p^2 + q^2)\frac{p^2q^2}{(p^2 + q^2)^2}}$$

asymptotically, where $q = 1 - p$ and p is the probability for the biased coin being H.

Generally, if the maximum depth of the tree is d , then we can calculate the efficiency of the random-stream algorithm by iteration in the following way:

Theorem 5. *When the maximum depth of the tree is d and the probability of the biased coin is p of being H, the expected number of coin tosses required per random bit is*

$$\frac{1}{\rho_d(p)}$$

asymptotically, where $\rho_d(p)$ can be obtained by iterating

$$\rho_d(p) = pq + \frac{1}{2}\rho_{d-1}(p^2 + q^2) + \frac{1}{2}(p^2 + q^2)\rho_{d-1}\left(\frac{p^2}{p^2 + q^2}\right) \quad (1)$$

with $q = 1 - p$ and $\rho_0(p) = pq$.

Theorem 5 shows that the efficiency of a random-stream algorithm with maximum depth d can be easily calculated by iteration. One thing that we can claim is,

$$\lim_{d \rightarrow \infty} \rho_d(p) = H(p).$$

However, it is difficult to get an explicit expression for $\rho_d(p)$ when d is finite. As d increases, the convergence rate of $\rho_d(p)$ depends on the value of p . The following extreme case implies that $\rho_d(p)$ can converge to $H(p)$ very quickly.

TABLE I
THE EXPECTED NUMBER OF COIN TOSSES REQUIRED PER RANDOM BIT FOR DIFFERENT PROBABILITY p AND DIFFERENT MAXIMUM DEPTHS

maximum depth	p=0.1	p=0.2	p=0.3	p=0.4	p=0.5
0	11.1111	6.2500	4.7619	4.1667	4.0000
1	5.9263	3.4768	2.7040	2.3799	2.2857
2	4.2857	2.5816	2.0299	1.7990	1.7297
3	3.5102	2.1484	1.7061	1.5190	1.4629
4	3.0655	1.9023	1.5207	1.3596	1.3111
5	2.7876	1.7480	1.4047	1.2598	1.2165
7	2.4764	1.5745	1.2748	1.1485	1.1113
10	2.2732	1.4619	1.1910	1.0772	1.0441
15	2.1662	1.4033	1.1478	1.0408	1.0101
∞	2.1322	1.3852	1.1347	1.0299	1.0000

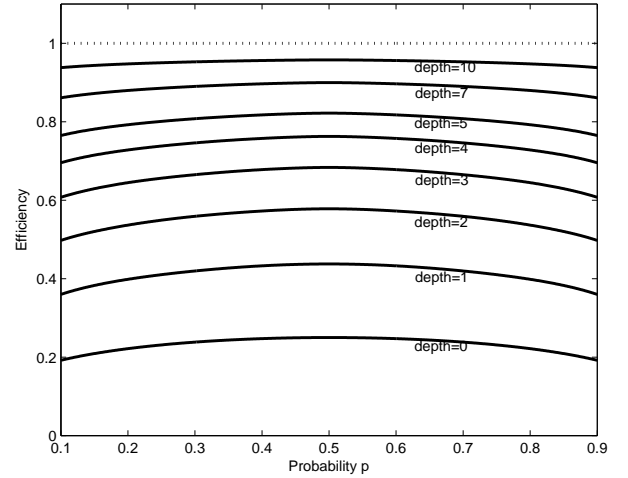


Fig. 2. The efficiency for different probability p and different maximum depths.

Example 5. *Let us consider the case that $p = \frac{1}{2}$. According to Equ. (1), we have*

$$\rho_d\left(\frac{1}{2}\right) = \frac{1}{4} + \frac{1}{2}\rho_{d-1}\left(\frac{1}{2}\right) + \frac{1}{4}\rho_{d-1}\left(\frac{1}{2}\right),$$

where $\rho_0(\frac{1}{2}) = \frac{1}{4}$. Based on this iterative relation, it can be obtained that

$$\rho_d\left(\frac{1}{2}\right) = 1 - \left(\frac{3}{4}\right)^{d+1}.$$

So when $p = \frac{1}{2}$, $\rho_d(p)$ can converge to $H(p) = 1$ very quickly as d increases.

In Table I, we tabulate the expected number of coin tosses required per random bit in the random-stream algorithm with different maximum depths. We see that as the maximum depth increases, the efficiency of the random-stream algorithm approaches the theoretical limitation quickly. Let us consider the case of $p = 0.3$ as an example. If the maximum depth is 0, the random-stream algorithm is as efficient as von Neumann's scheme, which requires expected 4.76 coin tosses to generate one random bit. If the maximum depth is 7, it requires only

TABLE II
THE EXPECTED TIME FOR PROCESSING A SINGLE INPUT COIN TOSS FOR
DIFFERENT PROBABILITY p AND DIFFERENT MAXIMUM DEPTHS

maximum depth	p=0.1	p=0.2	p=0.3	p=0.4	p=0.5
0	1.0000	1.0000	1.0000	1.0000	1.0000
1	1.9100	1.8400	1.7900	1.7600	1.7500
2	2.7413	2.5524	2.4202	2.3398	2.3125
3	3.5079	3.1650	2.9275	2.7840	2.7344
4	4.2230	3.6996	3.3414	3.1256	3.0508
5	4.8968	4.1739	3.6838	3.3901	3.2881
7	6.1540	4.9940	4.2188	3.7587	3.5995
10	7.9002	6.0309	4.8001	4.0783	3.8311
15	10.6458	7.5383	5.5215	4.3539	3.9599

expected 1.27 coin tosses to generate one random bit. That is very close to the theoretical limitation 1.13. However, the space cost of the algorithm has an exponential dependence on the maximum depth. That requires us to balance the efficiency and the space cost in real applications. Specifically, if we define efficiency as the ratio between the theoretical lower bound and the real value of the expected number of coin tosses, then Fig. 2 shows the relation between the efficiency and the maximum depth for different probability p .

Another property that we consider is the expected time for processing a single coin toss. Assume that it takes a single unit of time to process a message received at a node, then the expected time is exactly the expected number of messages that have been generated in the status tree (including the input coin toss itself). Table II shows the expected time for processing a single input bit when the input is infinitely long, implying the computational efficiency of the random-stream algorithm with limited depth. It can be proved that for an input generated by an arbitrary biased coin the expected time for processing a single coin toss is upper bounded by the maximum depth plus one (it is not a tight bound).

B. Proof of Theorem 2

In this subsection, we prove Theorem 2.

Lemma 6. Let \mathcal{T} be the status tree induced by $X_A \in \{H, T\}^*$, and let $k_1, k_2, \dots, k_{|\mathcal{T}|}$ be the number of bits generated by the nodes in \mathcal{T} , where $|\mathcal{T}|$ is the number of nodes in \mathcal{T} . Then for any $y_i \in \{0, 1\}^{k_i}$ with $1 \leq i \leq |\mathcal{T}|$, there exists a unique sequence $X_B \in \{H, T\}^*$ such that it induces the same status tree \mathcal{T} , and the bits generated by the i th node in \mathcal{T} is y_i . For such a sequence X_B , it is a permutation of X_A with the same last element.

Proof: To prove this conclusion, we can apply the idea of Lemma 1. It is obviously that if the maximum depth of \mathcal{T} is zero, then the conclusion is trivial. Assume that the conclusion holds for any status tree with maximum depth at most k , then we show that it also holds for any status tree with maximum depth $k + 1$.

Given a status tree \mathcal{T} with maximum depth $k + 1$, we use $Y_A \in \{0, 1\}^*$ to denote the binary sequence generated by the

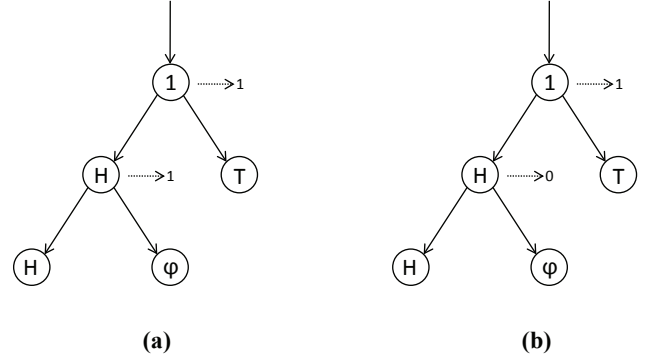


Fig. 3. An example for demonstrating Lemma 6, where the input sequence for (a) is HTTTHT, and the input sequence for (b) is TTHTHT.

root node based on X_A , and L_A, R_A to denote the sequences of symbols received by its left child and right child. According to our assumption, by flipping the bits generated by the left subtree, we can construct a unique sequence $L_B \in \{H, T\}^*$ uniquely such that L_B is a permutation of L_A with the same last element. Similarly, for the right subtree, we have $R_B \in \{H, T\}^*$ uniquely such that R_B is a permutation of R_A with the same last element.

Assume that by flipping the bits generated by the root node, we get a binary sequence Y_B such that $|Y_B| = |Y_A|$ (If the label $\alpha \in \{0, 1\}$, we add it to Y_A and Y_B), then

$$|L_B| = |Y_B| + |R_B|,$$

which implies that we can construct X_B from L_B, R_B, Y_B and the label α on the root node uniquely (according to the proof of the above lemma). Since the length of X_B is uniquely determined by $|L_B|$ and α , we can also conclude that X_A and X_B have the same length.

To see that X_B is a permutation of X_A , we show that X_B has the same number of H's as X_A . Given a sequence $X \in \{H, T\}^*$, let $w_H(X)$ denote the number of H's in X . It is not hard to see that

$$w_H(X_A) = w_H(L_A) + w_H(R_A) + w_H(\alpha),$$

$$w_H(X_B) = w_H(L_B) + w_H(R_B) + w_H(\alpha),$$

where $w_H(L_A) = w_H(L_B)$ and $w_H(R_A) = w_H(R_B)$ due to our assumption. Hence $w_H(X_A) = w_H(X_B)$ and X_B is a permutation of X_A .

Finally, we would like to see that X_A and X_B have the same last element. That is because if $\alpha \in \{H, T\}$, then both X_A and X_B end with α . If $\alpha \in \{\phi, 0, 1\}$, the last element of X_B depends on the last element of L_B , the last element of R_B , and α . Our assumption gives that L_B has the same element as L_A and R_B has the same element as R_A . So we can conclude that X_A and X_B have the same last element.

This completes the proof. \blacksquare

Example 6. The status tree of a sequence HTTTHT is given by Fig. 3(a). If we flip the second bit 1 into 0, see Fig. 3(b), we can construct a sequence of coin tosses, which is TTHTHT.

Now, we define an equivalence relation on $\{H, T\}^*$.

Definition 1. Let \mathcal{T}_A be the status tree of X_A and \mathcal{T}_B be the status tree of X_B . Two sequences $X_A, X_B \in \{H, T\}^*$ are equivalent denoted by $X_A \equiv X_B$ if and only if $\mathcal{T}_A = \mathcal{T}_B$, and for each pair of nodes (u, v) with $u \in \mathcal{T}_A$ and $v \in \mathcal{T}_B$ at the same position they generate the same number of bits.

Let S_Y with $Y \in \{0, 1\}^k$ denote the set consisting of all the binary sequences yielding Y . Here, we say that a binary sequence X yields Y if and only if $X[1 : |X| - 1]$ generates a sequence shorter than Y and X generates a sequence with Y as a prefix (including Y itself). Namely, let f be the function of the random-stream algorithm, then

$$|f(X[1 : |X| - 1])| < |Y|,$$

$$f(X) = Y\Delta \text{ with } \Delta \in \{0, 1\}^*.$$

To prove that the algorithm can generate random-bit streams, we show that for any distinct binary sequences $Y_1, Y_2 \in \{0, 1\}^k$, the elements in S_{Y_1} and those in S_{Y_2} are one-to-one mapping.

Lemma 7. Let f be the function of the random-stream algorithm. For any distinct binary sequences $Y_1, Y_2 \in \{0, 1\}^k$, if $X_A \in S_{Y_1}$, there are exactly one sequence $X_B \in S_{Y_2}$ such that

- $X_B \equiv X_A$.
- $f(X_A) = Y_1\Delta$ and $f(X_B) = Y_2\Delta$ for some binary sequence $\Delta \in \{0, 1\}^*$.

Proof: Let us prove this conclusion by induction. Here, we use X'_A to denote the prefix of X_A of length $|X_A| - 1$ and use β to denote the last symbol of X_A . So $X_A = X'_A\beta$.

When $k = 1$, if $X_A \in S_0$, we can write $f(X_A)$ as 0Δ for some $\Delta \in \{0, 1\}^*$. In this case, we assume that the status tree of X'_A is \mathcal{T}'_A , and in which node u generates the first bit 0 when reading the symbol β . If we flip the label of u from 0 to 1, we get another status tree, denoted by \mathcal{T}'_B . Using the same argument as Lemma 1, we are able to construct a sequence X'_B such that its status tree is \mathcal{T}'_B and it does not generate any bits. Concatenating X'_B with β results in a new sequence X_B , i.e., $X_B = X'_B\beta$, such that $X_B \equiv X_A$ and $f(X_B) = 1\Delta$. Similarly, for any sequence X_B that yields 1, i.e., $X_B \in S_1$, if $f(X_B) = 1\Delta$, we can find a sequence $X_A \in S_0$ such that $X_A \equiv X_B$ and $f(X_A) = 0\Delta$. So we can say that the elements in S_0 and S_1 are one-to-one mapping.

Now we assume that all the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping for all $Y_1, Y_2 \in \{0, 1\}^k$, then we show that this conclusion also holds for any $Y_1, Y_2 \in \{0, 1\}^{k+1}$. Two cases need to be considered.

1) Y_1, Y_2 end with the same bit. Without loss of generality, we assume this bit is 0, then we can write $Y_1 = Y'_10$ and $Y_2 = Y'_20$. If $X_A \in S_{Y'_1}$, then we can write $f(X_A) = Y'_1\Delta'$ in which the first bit of Δ' is 0. According to our assumption, there exists a sequence $X_B \in S_{Y'_2}$ such that $X_B \equiv X_A$ and $f(X_B) = Y'_2\Delta'$. In this case, if we write $f(X_A) = Y_1\Delta = Y'_10\Delta$, then $f(X_B) = Y'_2\Delta' = Y'_20\Delta = Y_2\Delta$. So such a sequence X_B satisfies our requirements.

If $X_A \notin S_{Y'_1}$, that means Y'_1 has been generated before reading the symbol β . Let us consider a prefix of X_A , denote

by $\overline{X_A}$, such that it yields Y'_1 . In this case, $f(X'_A) = Y'_1$ and we can write $X_A = \overline{X_A}Z$. According to our assumption, there exists exactly one sequence $\overline{X_B}$ such that $\overline{X_B} \equiv \overline{X_A}$ and $f(X'_B) = Y'_2$. Since $\overline{X_A}$ and $\overline{X_B}$ induce the same status tree, if we construct a sequence $X_B = \overline{X_B}Z$, then $X_B \equiv X_A$ and X_B generates the same bits as X_A when reading symbols from Z . It is easy to see that such a sequence X_B satisfies our requirements.

Since this result is also true for the inverse case, if Y_1, Y_2 end with same bit the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping.

2) Let us consider the case that Y_1, Y_2 end with different bits. Without loss of generality, we assume that $Y_1 = Y'_10$ and $Y_2 = Y'_21$. According to the argument above, the elements in $S_{00\dots00}$ and $S_{Y'_10}$ are one-to-one mapping; and the elements in $S_{00\dots01}$ and $S_{Y'_21}$ are one-to-one mapping. So our task is to prove that the elements in $S_{00\dots00}$ and $S_{00\dots01}$ are one-to-one mapping. For any sequence $X_A \in S_{00\dots00}$, let X'_A be its prefix of length $|X_A| - 1$. Here, X'_A generates only zeros whose length is at most k . Let \mathcal{T}'_A denote the status tree of X'_A and let u be the node in \mathcal{T}'_A that generates the $k + 1$ th bit (zero) when reading the symbol β . Then we can construct a new sequence X'_B with status tree \mathcal{T}'_B such that

- \mathcal{T}'_B and \mathcal{T}'_A are the same except the label of u is 0 and the label of the node at the same position in \mathcal{T}'_B is 1.
- For each node u in \mathcal{T}'_A , let v be its corresponding node at the same position in \mathcal{T}'_B , then u and v generate the same bits.

The construction of X'_B follows the proof of Lemma 6. If we construct a sequence $X_B = X'_B\beta$, it is not hard to show that X_B satisfies our requirements, i.e.,

- $X_B \equiv X_A$;
- X'_B generates less than $k + 1$ bits, i.e., $|f(X'_B)| \leq k$;
- If $f(X_A) = 0^k0\Delta$, then $f(X_B) = 0^k1\Delta$, where 0^k is for k zeros.

Also based on the inverse argument, we see that the elements in $S_{00\dots00}$ and $S_{00\dots01}$ are one-to-one mapping. So if Y_1, Y_2 end with different bits, the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping.

Finally, we can conclude that the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping for any $Y_1, Y_2 \in \{0, 1\}^k$ with $k > 0$.

This completes the proof. \blacksquare

Theorem 2. Given a source of biased coin with unknown probability, the random-stream algorithm generates a stream of random bits, i.e., for any $k > 0$, if we stop running the algorithm after generating k bits then these k bits are independent and unbiased.

Proof: According to Lemma 7, for any $Y_1, Y_2 \in \{0, 1\}^k$, the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping. If two sequences are one-to-one mapping, they have to be equivalent, which implies that their probabilities of being generated are the same. Hence, the probability of generating a sequence in S_{Y_1} is equal to that of generating a sequence in S_{Y_2} . It implies that Y_1 and Y_2 have the same probability of being generated for a fixed number k . Since this is true for any $Y_1, Y_2 \in \{0, 1\}^k$, the probability of generating an arbitrary binary sequence $Y \in \{0, 1\}^k$ is 2^{-k} . Finally, we have the statement in the theorem.

This completes the proof. \blacksquare

C. Proof of Theorem 3

Lemma 8. *Given a stream of biased coin tosses, where the probability of generating H is p , we run the random-stream algorithm until the number of coin tosses reaches l . In this case, let m be the number of random bits generated, then for any $\epsilon, \delta > 0$, if l is large enough, we have that*

$$P\left[\frac{m - lH(p)}{lH(p)} < -\epsilon\right] < \delta,$$

where

$$H(p) = -p \log_2 p - (1-p) \log_2 (1-p)$$

is the entropy of the biased coin.

Proof: If we consider the case of fixed input length, then the random-stream algorithm is as efficient as Peres's algorithm asymptotically. Using the same proof given in [5] for Peres's algorithm, we can get

$$\lim_{l \rightarrow \infty} \frac{E[m]}{l} = H(p).$$

Given a sequence of coin tosses of length l , we want to prove that for any $\epsilon > 0$,

$$\lim_{l \rightarrow \infty} P\left[\frac{m - E[m]}{E[m]} < -\epsilon\right] = 0.$$

To prove this result, we assume that this limitation holds for $\epsilon = \epsilon_1$, i.e., for any $\delta > 0$, if l is large enough, then

$$P\left[\frac{m - E[m]}{E[m]} < -\epsilon_1\right] < \delta.$$

Under this assumption, we show that there always exists $\epsilon_2 < \epsilon_1$ such that the limitation also holds for $\epsilon = \epsilon_2$. Hence, the value of ϵ can be arbitrarily small.

In the random-stream algorithm, l is the number of symbols (coin tosses) received by the root. Let m_1 be the number of random bits generated by the root, $m_{(l)}$ be the number of random bits generated by its left subtree and $m_{(r)}$ be the number of random bits generated by its right subtree. Then it is easy to get

$$m = m_1 + m_{(l)} + m_{(r)}.$$

Since the m_1 random bits generated by the root node are independent, we can always make l large enough such that

$$P\left[\frac{m_1 - E[m_1]}{E[m_1]} < -\epsilon_1/2\right] < \delta/3.$$

At the same time, by making l large enough, we can make both $m_{(l)}$ and $m_{(r)}$ large enough such that (based on our assumption)

$$P\left[\frac{m_{(l)} - E[m_{(l)}]}{E[m_{(l)}]} < -\epsilon_1\right] < \delta/3$$

and

$$P\left[\frac{m_{(r)} - E[m_{(r)}]}{E[m_{(r)}]} < -\epsilon_1\right] < \delta/3.$$

Based on the three inequalities above, we can get

$$P[m - E[m] \leq -\epsilon_1 \left(\frac{E[m_1]}{2} + E[m_{(l)}] + E[m_{(r)}] \right)] < \delta.$$

If we set

$$\epsilon_2 = \epsilon_1 \frac{\frac{E[m_1]}{2} + E[m_{(l)}] + E[m_{(r)}]}{E[m_1 + m_{(l)} + m_{(r)}]},$$

then

$$P\left[\frac{m - E[m]}{E[m]} < -\epsilon_2\right] < \delta.$$

Given the probability p of the coin, when l is large,

$$E[m_1] = \Theta(E[m]), E[m_{(l)}] = \Theta(E[m]), E[m_{(r)}] = \Theta(E[m]),$$

which implies that $\epsilon_2 < \epsilon_1$.

So we can conclude that for any $\epsilon > 0, \delta > 0$, if l is large enough then

$$P\left[\frac{m - E[m]}{E[m]} < -\epsilon\right] < \delta.$$

And based on the fact that $E[m] \rightarrow lH(p)$, we get the result in the lemma. \blacksquare

Theorem 3. *Given a biased coin with probability p being H , let n be the number of coin tosses required for generating k random bits in the random-stream algorithm, then*

$$\lim_{k \rightarrow \infty} \frac{E[n]}{k} = \frac{1}{H(p)}.$$

Proof: For any $\epsilon, \delta > 0$, we set $l = \frac{k}{H(p)}(1+\epsilon)$, according to the conclusion of the previous lemma, with probability at least $1-\delta$, the output length is at least k if the input length l is fixed and large enough. In another word, if the output length is k , which is fixed, then with probability at least $1-\delta$, the input length $n \leq l$.

So with probability less than δ , we require more than l coin tosses. The worst case is that we did not generate any bits for the first l coin tosses. In this case, we need to generate k more random bits. As a result, the expected number of coin tosses required is at most $l + E[n]$.

Based on the analysis above, we derive

$$E[n] \leq (1-\delta)l + (\delta)(l + E[n]),$$

then

$$E[n] \leq \frac{l}{1-\delta} = \frac{k}{H(p)} \frac{(1+\epsilon)}{(1-\delta)}.$$

Since ϵ, δ can be arbitrarily small when l (or k) is large enough

$$\lim_{k \rightarrow \infty} \frac{E[n]}{k} \leq \frac{1}{H(p)}.$$

Based on Shannon's theory [2], it is impossible to generate k random bits from a source with expected entropy less than k . Hence

$$\lim_{k \rightarrow \infty} E[nH(p)] \geq k,$$

i.e.,

$$\lim_{k \rightarrow \infty} \frac{E[n]}{k} \geq \frac{1}{H(p)}.$$

Finally, we get the conclusion in the theorem. This completes the proof. \blacksquare

D. Proof of Theorem 4

The proof of Theorem 4 is very similar as the proof of Theorem 2. Let S_Y with $Y \in \{0, 1\}^k$ denote the set consisting of all the binary sequences yielding Y in the random-stream algorithm with limited maximum depth. Then for any distinct binary sequences $Y_1, Y_2 \in \{0, 1\}^k$, the elements in S_{Y_1} and those in S_{Y_2} are one-to-one mapping. Specifically, we can get the following lemma.

Lemma 9. *Let f be the function of the random-stream algorithm with maximum depth d . For any distinct binary sequences $Y_1, Y_2 \in \{0, 1\}^k$, if $X_A \in S_{Y_1}$, there exists one sequence $X_B \in S_{Y_2}$ such that*

- $X_A \equiv X_B$.
- Let \mathcal{T}_A be the status tree of X_A and \mathcal{T}_B be the status tree of X_B . For any node u with depth larger than d in \mathcal{T}_A , let v be its corresponding node in \mathcal{T}_B at the same position, then u and v generate the same bits.
- $f(X_A) = Y_1\Delta$ and $f(X_B) = Y_2\Delta$ for some binary sequence $\Delta \in \{0, 1\}^*$.

Proof: The proof of this lemma is a simple modification of that for Lemma 7, which is by induction. A simple sketch is given as follows.

First, similar as the proof for Lemma 7, it can be proved that: when $k = 1$, for any sequence $X_A \in S_0$, there exists one sequence $X_B \in S_1$ such that X_A, X_B satisfy the conditions in the lemma, and vice versa. So we can say that the elements in S_0 and S_1 are one-to-one mapping. Then we assume that all the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping for all $Y_1, Y_2 \in \{0, 1\}^k$, then we show that this conclusion also holds for any $Y_1, Y_2 \in \{0, 1\}^{k+1}$. Two cases need to be considered.

1) Y_1, Y_2 end with the same bit. Without loss of generality, we assume this bit is 0, then we can write $Y_1 = Y'_10$ and $Y_2 = Y'_20$.

If $X_A \in S_{Y'_1}$, then according to our assumption, it is easy to prove the conclusion, i.e., there exists a sequence X_B satisfies the conditions.

If $X_A \notin S_{Y'_1}$, then we can write $X_A = \overline{X_A}Z$ and $\overline{X_A} \in S_{Y'_1}$. According to our assumption, for the sequence $\overline{X_A}$, we can find its mapping $\overline{X_B}$ such that (1) $\overline{X_A} \equiv \overline{X_B}$; (2) $\overline{X_A}, \overline{X_B}$ induce the same status tree and their corresponding nodes with depth larger than d generate the same bits; and (3) $f(\overline{X_A}) = Y'_1$ and $f(\overline{X_B}) = Y'_2$. If we construct a sequence $X_B = \overline{X_B}Z$, it will satisfy all the conditions in the lemma.

Since this result is also true for the inverse case, if Y_1, Y_2 end with same bit, the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping.

2) Y_1, Y_2 end with different bits. Without loss of generality, we assume that $Y_1 = Y'_10$ and $Y_2 = Y'_21$. According to the argument above, the elements in S_{0^k0} and $S_{Y'_1}$ are one-to-one mapping; and the elements in S_{0^k1} and $S_{Y'_2}$ are one-to-one mapping. So we only need to prove that the elements in S_{0^k0} and S_{0^k1} are one-to-one mapping. In this case, for any $X_A \in S_{0^{k-1}0}$, let $X_A = X'_A\beta$ with a single symbol β . Then X'_A generates only zeros whose length is at most k . Let \mathcal{T}'_A denote the status tree of X'_A and let u be the node in \mathcal{T}'_A that generates the $k + 1$ th bit (zero) when reading the symbol β .

Note that the depth of u is at most d . In this case, we can construct a new sequence X'_B with status tree \mathcal{T}'_B such that

- \mathcal{T}'_B and \mathcal{T}'_A are the same except the label of u is 0 and the label of the node at the same position in \mathcal{T}'_B is 1.
- For each node u in \mathcal{T}'_A , let v be its corresponding node at the same position in \mathcal{T}'_B , then u and v generate the same bits.

Then we can prove that the sequence $X_B = X'_B\beta$ satisfies our all our conditions in the lemma. Also based on the inverse argument, we can claim that the elements in S_{0^k0} and S_{0^k1} are one-to-one mapping.

Finally, we can conclude that the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping for any $Y_1, Y_2 \in \{0, 1\}^k$ with $k > 0$. This completes the proof. ■

From the above lemma, it is easy to get Theorem 4.

Theorem 4. *Given a source of biased coin with unknown probability, the random-stream algorithm with maximum depth d generates a stream of random bits, i.e., for any $k > 0$, if we stop running the algorithm after generating k bits then these k bits are independent and unbiased.*

Proof: We can apply the same procedure of proving Theorem 3. ■

E. Proof of Theorem 5

Similar to the proof of Theorem 3, we first consider the case that the input length is fixed.

Lemma 10. *Given a stream of biased coin tosses, where the probability of generating H is p , we run the random-stream algorithm with maximum depth d until the number of coin tosses reaches l . In this case, let m be the number of random bits generated, then for any $\epsilon, \delta > 0$, if l is large enough, we have that*

$$P\left[\frac{m - l\rho_d(p)}{l\rho_d(p)} < -\epsilon\right] < \delta,$$

where $\rho_d(p)$ is given in (1).

Proof: Let $\rho_d(p)$ be the asymptotic expected number of random bits generated per coin toss when the random-stream algorithm has maximum depth d and the probability of the biased coin is p . Then

$$\lim_{l \rightarrow \infty} \frac{E[m]}{l} = \rho_d(p).$$

When the fixed input length l is large enough, the random-stream algorithm generates approximately $l\rho_d(p)$ random bits, which are generated by the root node, the left subtree (subtree rooted at root's left child) and the right subtree (subtree rooted at the root's right child). Considering the root node, it generates approximately lpq random bits with $q = 1 - p$. Meanwhile, the root node passes approximately $\frac{l}{2}$ messages (H or T) to its left child, where the messages are independent and the probability of H is $p^2 + q^2$; and the root node passes approximately $\frac{l}{2}(p^2 + q^2)$ messages (H or T) to its right child, where the messages are independent and the probability of H is $\frac{p^2}{p^2 + q^2}$. As a result, according to the definition of ρ_d , the left subtree generates approximately $\frac{l}{2}\rho_{d-1}(p^2 + q^2)$

random bits, and the right subtree generates approximately $\frac{l}{2}(p^2 + q^2)\rho_{d-1}(\frac{p^2}{p^2+q^2})$ random bits. As $l \rightarrow \infty$, we have

$$\lim_{l \rightarrow \infty} \frac{l\rho_d(p)}{lpq + \frac{l}{2}\rho_{d-1}(p^2 + q^2) + \frac{l}{2}(p^2 + q^2)\rho_{d-1}(\frac{p^2}{p^2+q^2})} = 1.$$

It yields

$$\rho_d(p) = pq + \frac{1}{2}\rho_{d-1}(p^2 + q^2) + \frac{1}{2}(p^2 + q^2)\rho_{d-1}(\frac{p^2}{p^2+q^2}).$$

So we can calculate $\rho_d(p)$ by iteration. When $d = 0$, the status tree has the single root node, and it is easy to get $\rho_0(p) = pq$.

Then, following the proof of Lemma 8, for any $\epsilon, \delta > 0$, if l is large enough, we have that

$$P[\frac{m - E[m]}{E[m]} < -\epsilon] < \delta.$$

So we can get the conclusion in the lemma. This completes the proof. ■

From the above lemma, we can get Theorem 5, that is,

Theorem 5. *When the maximum depth of the tree is d and the probability of the biased coin is p of being H, the expected number of coin tosses required per random bit is*

$$\frac{1}{\rho_d(p)}$$

asymptotically, where $\rho_d(p)$ can be obtained by iterating

$$\rho_d(p) = pq + \frac{1}{2}\rho_{d-1}(p^2 + q^2) + \frac{1}{2}(p^2 + q^2)\rho_{d-1}(\frac{p^2}{p^2+q^2}) \quad (1)$$

with $q = 1 - p$ and $\rho_0(p) = pq$.

Proof: We can apply the same procedure of proving Theorem 2 except we apply Lemma 10 instead of Lemma 8. ■

III. GENERALIZED RANDOM-STREAM ALGORITHM

A. Preliminary

In [10], we introduced a universal scheme for transforming an arbitrary algorithm for generating random bits from a sequence of biased coin tosses to manage the general source of an m -sided die. This scheme works when the input is a sequence of fixed length; in this section, we study how to modify this scheme to generate random-bit streams from m -sided dice. For sake of completeness we describe the original scheme here.

The main idea of the scheme is to convert a sequence with alphabet larger than two, written as

$$X = x_1x_2\dots x_n \in \{0, 1, \dots, m-1\}^n,$$

into multiple binary sequences. To do this, we create a binary tree, called a binarization tree, in which each node is labeled with a binary sequence of H and T. Given the binary representations of x_i for all $1 \leq i \leq n$, the path of each node in the tree indicates a prefix, and the binary sequence labeled at this node consists of all the bits (H or T) following the prefix in the binary representations of x_1, x_2, \dots, x_n (if it exists). Fig. 4 is an instance of binarization tree when the

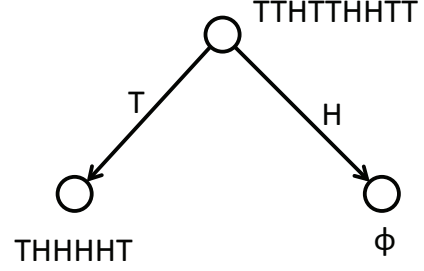


Fig. 4. An instance of binarization tree.

input sequence is $X = 012112210$, produced by a 3-sided die. To see this, we write each symbol (die roll) into a binary representation of length two, hence X can be represented as

$$TT, TH, HT, TH, TH, HT, HT, TH, TT.$$

Only collecting the first bits of all the symbols yields an independent binary sequence

$$X_\phi = TTHTTHHTT,$$

which is labeled on the root node; Collecting the second bits following T, we get another independent binary sequence

$$X_T = THHHHT,$$

which is labeled on the left child of the root node.

The universal scheme says that we can ‘treat’ each binary sequence labeled on the binarization tree as a sequence of biased coin tosses: Let Ψ be any algorithm that can generate random bits from an arbitrary biased coin, then applying Ψ to each of the sequences labeled on the binarization tree and concatenating their outputs together results in an independent and unbiased sequence, namely, a sequence of random bits.

Specifically, given the number of sides m of a loaded die, the depth of the binarization tree is $b = \lceil \log_2 m \rceil - 1$. Let Υ_b denote the set consisting of all the binary sequences of length at most b , i.e.,

$$\Upsilon_b = \{\phi, T, H, TT, TH, HT, HH, \dots, HHH\dots HH\}.$$

Given $X \in \{0, 1, \dots, m-1\}^n$, let X_γ denote the binary sequence labeled on a node corresponding to a prefix γ in the binarization tree, then we get a group of binary sequences

$$X_\phi, X_T, X_H, X_{TT}, X_{TH}, X_{HT}, X_{HH}, \dots$$

For any function Ψ that generates random bits from a fixed number of coin tosses, we can generate random bits from X by calculating

$$\Psi(X_\phi) + \Psi(X_T) + \Psi(X_H) + \Psi(X_{TT}) + \Psi(X_{TH}) + \dots,$$

where $A + B$ is the concatenation of A and B .

So in the above example, the output of $X = 012112210$ is $\Psi(X_\phi) + \Psi(X_T)$, i.e.,

$$\Psi(TTHTTHHTT) + \Psi(THHHHT).$$

This conclusion is simple, but not obvious, since the binary sequences labeled on the same binarization tree are correlated with each other.

B. Generalized Random-Stream Algorithm

We want to generalize the random-stream algorithm to generate random-bit streams from an m -sided die. Using the similar idea as above, we convert the input stream into multiple binary streams, where each binary stream corresponds to a node in the binalization tree. We apply the random-stream algorithm to all these binary streams individually, and for each stream we create a status tree for storing state information. When we read a dice roll of m sides from the source, we pass all the $\log_2 m$ bits of its binary representation to $\lceil \log_2 m \rceil$ different streams that corresponds to a path in the binalization tree. Then we process all these $\lceil \log_2 m \rceil$ streams from top to bottom along that path. In this way, a single binary stream is produced. While each node in the binalization tree generates a random-bit stream, the resulting single stream is a mixture of these random-bit streams. But it is not obvious whether the resulting stream is a random-bit stream or not, since the values of the bits generated affect their orders.

The following example is constructed for demonstrating this algorithm.

Let us consider a stream of symbols generated from a 3-sided die,

012112210...

Instead of storing a binary sequence at each node in the binalization tree, we associate each node with a status tree corresponding to a random-stream algorithm. Here, we get two nontrivial binary streams

TTHTTHHTT..., THHHHT...

corresponding to prefix ϕ and T respectively, Fig. 5 demonstrates how the status trees change when we read symbols one by one. For instance, when the 4th symbol 1(TH) is read, it passes T to the root node (corresponding to the prefix ϕ) and passes H to the left child of the root node (corresponding to the prefix T) of the binalization tree. Based on the rules of the random-stream algorithm, we modify the status trees associated with these two nodes. During this process, a bit 0 is generated.

Finally, this scheme generates a stream of bits 010..., where the first bit is generated after reading the 4th symbol, the second bit is generated after reading the 5th symbol, ... We call this scheme as the generalized random-stream algorithm. As we expected, this algorithm can generate a stream of random bits from an arbitrary loaded die with $m \geq 2$ sides.

Theorem 11. *Given a loaded die with $m \geq 2$ sides, if we stop running the generalized random-stream algorithm after generating k bits, then these k bits are independent and unbiased.*

The proof of the above theorem is given in Subsection III-C.

Since the random-stream algorithm is as efficient as Peres's algorithm asymptotically, we can prove that the generalized random-stream algorithm is also asymptotically optimal.

Theorem 12. *Given an m -sided die with probability distribution $\rho = (p_0, p_1, \dots, p_{m-1})$, let n be the number of symbols*

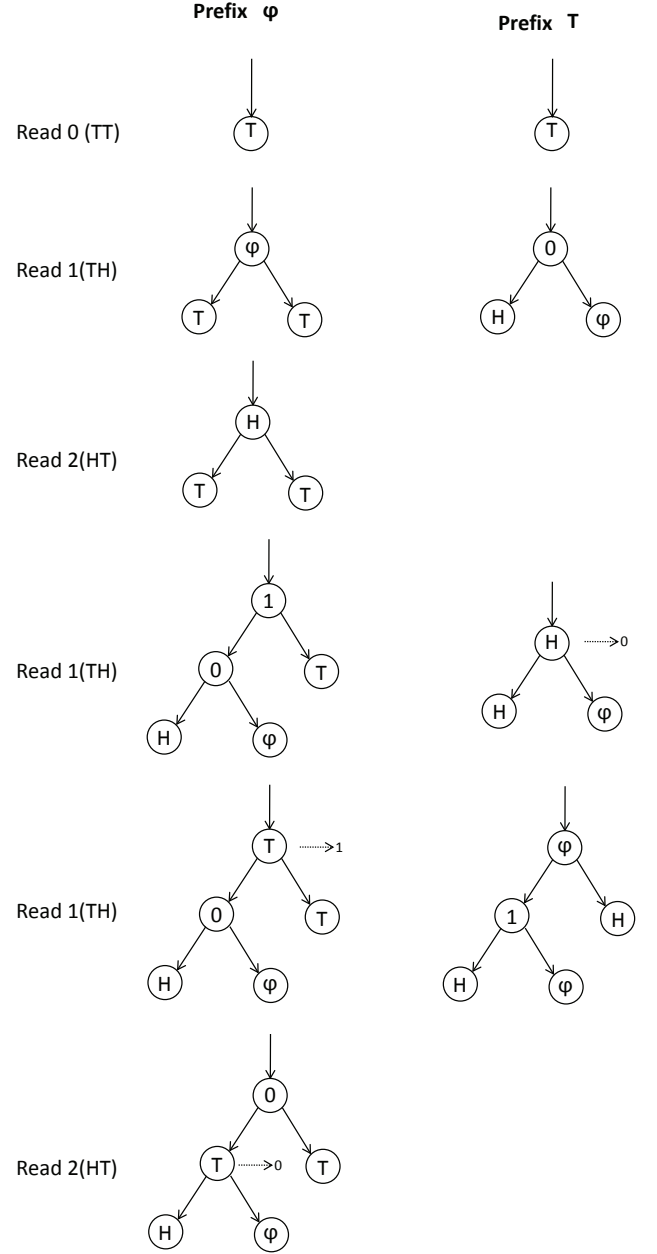


Fig. 5. The changes of status trees in the generalized random-stream algorithm when the input stream is 012112210....

(dice rolls) used in the generalized random-stream algorithm and let k be the number of random bits generated, then

$$\lim_{k \rightarrow \infty} \frac{E[n]}{k} = \frac{1}{H(p_0, p_1, \dots, p_{m-1})},$$

where

$$H(p_0, p_1, \dots, p_{m-1}) = \sum_{i=0}^{m-1} p_i \log_2 \frac{1}{p_i}$$

is the entropy of the m -sided die.

Proof: First, according to Shannon's theory, it is easy to

get that

$$\lim_{k \rightarrow \infty} \frac{E[n]}{k} \geq \frac{1}{H(p_0, p_1, \dots, p_{m-1})}.$$

Now, we let

$$n = \frac{k}{H(p_0, p_1, \dots, p_{m-1})} (1 + \epsilon)$$

with an arbitrary $\epsilon > 0$. Following the proof of Theorem 7 in [10], it can be shown that when k is large enough, the algorithm generates more than k random bits with probability at least $1 - \delta$ with any $\delta > 0$. Then using the same argument in Theorem 3, we can get

$$\lim_{k \rightarrow \infty} \frac{E[n]}{k} \leq \frac{1}{H(p_0, p_1, \dots, p_{m-1})} \frac{1 + \epsilon}{1 - \delta},$$

for any $\epsilon, \delta > 0$.

Hence, we can get the conclusion in the theorem. ■

Of source, we can limit the depths of all the status trees for saving space, with proof omitted. It can be seen that given a loaded die of m sides, the space usage is proportional to m and the expected computational time is proportional to $\log m$.

C. Proof of Theorem 11

Here, we want to prove that the generalized random-stream algorithm generates a stream of random bits from an arbitrary m -sided die. Similar as above, we let S_Y with $Y \in \{0, 1\}^k$ denote the set consisting of all the sequences yielding Y . Here, we say that a sequence X yields Y if and only if $X[1 : |X| - 1]$ generates a sequence shorter than Y and X generates a sequence with Y as a prefix (including Y itself). We would like to show that the elements in S_{Y_1} and those in S_{Y_2} are one-to-one mapping if Y_1 and Y_2 have the same length.

Definition 2. Two sequences $X_A, X_B \in \{0, 1, \dots, m-1\}^*$ with $m > 2$ are equivalent, denoted by $X_A \equiv X_B$, if and only $X_\gamma^A \equiv X_\gamma^B$ for all $\gamma \in \Upsilon_b$, where X_γ^A is the binary sequence labeled on a node corresponding to a prefix γ in the binalization tree induced by X_A , and the equivalence of X_γ^A and X_γ^B was given in Definition 1.

Lemma 13. [10] Let $\{X_\gamma^A\}$ with $\gamma \in \Upsilon_b$ be the binary sequences labeled on the binalization tree of $X_A \in \{0, 1, \dots, m-1\}^n$ as defined above. Assume X_γ^B is a permutation of X_γ^A for all $\gamma \in \Upsilon_b$, then there exists exactly one sequence $X_B \in \{0, 1, \dots, m-1\}^n$ such that it yields a binalization tree that labels $\{X_\gamma^B\}$ with $\gamma \in \Upsilon_b$.

Proof: The proof is provided in [10]. ■

Lemma 14. Let f be the function of the generalized random-stream algorithm, and let X_A be a sequence produced by an m -sided die. For any distinct sequences $Y_1, Y_2 \in \{0, 1\}^k$, if $X_A \in S_{Y_1}$, there are exactly one sequence $X_B \in S_{Y_2}$ such that

- $X_B \equiv X_A$.
- $f(X_A) = Y_1\Delta$ and $f(X_B) = Y_2\Delta$ for some binary sequence Δ .

Proof: The idea of the proof is to combine the proof of Lemma 7 with the result in Lemma 13.

Let us prove this conclusion by induction. Here, we use X'_A to denote the prefix of X_A of length $|X_A| - 1$ and use β to denote the last symbol of X_A . So $X_A = X'_A\beta$. X_γ^A is the binary sequence labeled on a node corresponding to a prefix γ in the binalization tree induced by X'_A , and the status tree of $X_\gamma^{A'}$ with $\gamma \in \Upsilon_b$ is denoted as \mathcal{T}_γ^A .

When $k = 1$, if $X_A \in S_0$, we can write $f(X_A)$ as 0Δ . In this case, let u in \mathcal{T}_θ^A with $\theta \in \Upsilon_b$ be the node that generates the first bit 0. If we flip the label of u from 0 to 1, we get another status tree \mathcal{T}_θ^B . Using the same argument in Lemma 6, we are able to construct a sequence X_θ^B such that its status tree is \mathcal{T}_θ^B and it does not generate any bits. Here, X_θ^B is a permutation of X_θ^A . From $X_\theta^A, X_\theta^A, \dots, X_\theta^B, \dots$, we can construct a sequence X_B' uniquely following the proof of Lemma 13 (see [10]). Concatenating X_B' with β results in a new sequence X_B , i.e., $X_B = X_B'\beta$ such that $X_B \equiv X_A$ and $f(X_B) = 1\Delta$. Inversely, we can get the same result. It shows that the elements in S_0 and S_1 are one-to-one mapping.

Now we assume that the conclusion holds for all $Y_1, Y_2 \in \{0, 1\}^k$, then we show that it also holds for any $Y_1, Y_2 \in \{0, 1\}^{k+1}$. Two cases need to be considered.

1) Y_1, Y_2 end with the same bit. Without loss of generality, we assume that this bit is 0, then we can write $Y_1 = Y_1'0$ and $Y_2 = Y_2'0$. If X_A yields Y_1 , based on our assumption, it is easy to see that there exists a sequence X_B satisfies our requirements. If X_A does not yield Y_1 , that means Y_1' has been generated before reading the symbol β . Let us consider a prefix of X_A , denote by \bar{X}_A , such that it yields Y_1' . In this case, $f(X_A') = Y_1'$ and we can write $X_A = X_A'Z$. According to our assumption, there exists exactly one sequence \bar{X}_B such that $\bar{X}_B \equiv \bar{X}_A$ and $f(X_B') = Y_2'$. Since \bar{X}_A and \bar{X}_B lead to the same binalization tree (all the status trees at the same positions are the same), if we construct a sequence $X_B = \bar{X}_BZ$, then $X_B \equiv X_A$ and X_B generates the same bits as X_A when reading symbols from Z . It is easy to see that such a sequence X_B satisfies our requirements.

Since this result is also true for the inverse case, if $Y_1, Y_2 \in \{0, 1\}^{k+1}$ end with the same bit, the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping.

2) Let us consider the case that Y_1, Y_2 end with different bits. Without loss of generality, we assume that $Y_1 = Y_1'0$ and $Y_2 = Y_2'1$. According to the argument above, the elements in $S_{00\dots00}$ and $S_{Y_1'0}$ are one-to-one mapping; the elements in $S_{00\dots01}$ and $S_{Y_2'1}$ are one-to-one mapping. So our task is to prove that the elements in $S_{00\dots00}$ and $S_{00\dots01}$ are one-to-one mapping. For any sequence $X_A \in S_{00\dots00}$, let X'_A be its prefix of length $|X_A| - 1$. Here, X'_A generates only zeros whose length is at most k . Let \mathcal{T}_θ^A denote one of the status trees such that $u \in \mathcal{T}_\theta^A$ is the node that generates that $k+1$ th bit (zero) when reading the symbol β . Then we can construct a new sequence X_B' such that

- Let $\{X_\gamma^B\}$ with $\gamma \in \Upsilon_b$ be the binary sequences induced by X_B' , and let \mathcal{T}_γ^B be the status tree of X_γ^B . The binalization trees of X'_A and X_B' are the same (all the status trees at the same positions are the same), except the label of u is 0 and the label of its corresponding node v in \mathcal{T}_θ^B is 1.

- Each node u in \mathcal{T}_γ^B generates the same bits as its corresponding node v in \mathcal{T}_γ^A for all $\gamma \in \Upsilon_b$.

The construction of X'_B follows the proof of Lemma 1 and then Lemma 13. If we construct a sequence $X_B = X'_B\beta$, it is not hard to show that X_B satisfies our requirements, i.e.,

- $X_B \equiv X_A$;
- X'_B generates less than $k + 1$ bits, i.e., $|f(X'_B)| \leq k$;
- If $f(X_A) = Y_1\Delta = Y'_10\Delta$, then $f(X_B) = Y'_11\Delta = Y_2\Delta$.

Also based on the inverse argument, we see that the elements in $S_{00...00}$ and $S_{00...01}$ are one-to-one mapping.

Finally, we can conclude that the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping for any $Y_1, Y_2 \in \{0, 1\}^k$ with $k > 0$.

This completes the proof. ■

Based on the above result and the argument for Theorem 2, we can easily prove Theorem 11.

Theorem 11. *Given a loaded die with $m \geq 2$ sides, if we stop running the generalized random-stream algorithm after generating k bits, then these k bits are independent and unbiased.*

IV. EXTENSION FOR MARKOV CHAINS

In this section, we study how to efficiently generate random-bit streams from Markov chains. The nonstream case was studied by Samuelson [7], Blum [1] and later generalized by Zhou and Bruck [9]. Here, using the techniques developed in [9], and applying the techniques introduced in this paper, we are able to generate random-bit streams from Markov chains. We present the algorithm briefly.

For a given Markov chain, it generates a stream of states, denoted by $x_1x_2x_3... \in \{s_1, s_2, ..., s_m\}^*$. We can treat each state, say s , as a die and consider the ‘next states’ (the states the chain has transitioned to after being at state s) as the results of a die roll, called the exit of s . For all $s \in \{s_1, s_2, ..., s_m\}$, if we only consider the exits of s , they form a stream. So we have total m streams corresponding to the exits of $s_1, s_2, ..., s_m$ respectively. For example, assume the input is

$$X = s_1s_4s_2s_1s_3s_2s_3s_1s_1s_2s_3s_4s_1...$$

If we consider the states following s_1 , we get a stream as the set of states in boldface:

$$X = s_1\mathbf{s_4}s_2s_1\mathbf{s_3}s_2s_3s_1\mathbf{s_1s_2}s_3s_4s_1...$$

Hence the four streams are

$$s_4s_3s_1s_2..., s_1s_3s_3..., s_2s_1s_4..., s_2s_1...$$

The generalized random-stream algorithm is applied to each stream separately for generating random-bit streams. Here, when we get an exit of a state s , we should not directly pass it to the generalized random-stream algorithm that corresponds to the state s . Instead, it waits until we get the next exit of the state s . In another word, we keep the current exit in pending. In the above example, after we read $s_1s_4s_2s_1s_3s_2s_3s_1s_1s_2s_3s_4s_1$, $s_4s_3s_1$ has been passed to the generalized random-stream algorithm corresponding to s_1 , s_1s_3 has been passed to the generalized random-stream

algorithm corresponding to $s_2, ...,$ the most recent exit of each state, namely s_2, s_3, s_4, s_1 are in pending. Finally, we mix all the bits generated from different streams based on their natural generating order. As a result, we get a stream of random bits from an arbitrary Markov chain, and it achieves the information-theoretic upper bound on efficiency.

Now, we call this algorithm the random-stream algorithm for Markov chains, and we describe it as follows.

Input: A stream $X = x_1x_2x_3...$ produced by a Markov chain, where $x_i \in S = \{s_1, s_2, ..., s_m\}$.

Output: A stream of 0's and 1's.

Main Function:

Let Φ_i be the generalized random-stream algorithm for the exits of s_i for $1 \leq i \leq m$, and θ_i be the pending exit of s_i for $1 \leq i \leq m$.

Set $\theta_i = \phi$ for $1 \leq i \leq m$.

for each symbol x_j read from the Markov chain **do**

if $x_{j-1} = s_i$ **then**

if $\theta_i \neq \phi$ **then**

Input θ_i to Φ_i for processing.

end if

Set $\theta_i = x_j$.

end if

end for

Theorem 15. *Given a source of a Markov chain with unknown transition probabilities, the random-stream algorithm for Markov chains generates a stream of random bits, i.e., for any $k > 0$, if we stop running the algorithm after generating k bits then these k bits are independent and unbiased.*

The proof of the above theorem is a simple extension of the proof for Theorem 11. Let S_Y denote the set of input sequences that yield a binary sequence Y . Our main idea is still to prove that all the elements in S_{Y_1} and S_{Y_2} are one-to-one mapping for all $Y_1, Y_2 \in \{0, 1\}^k$ with $k > 0$. The detailed proof is a little complex, but it is not difficult; we only need to follow the proof of Theorem 11 and combine it with the following result from [9]. Here, we omit the detailed proof.

Lemma 16. *Given an input sequence $X = x_1x_2...x_N \in \{s_1, s_2, ..., s_m\}^N$ that produced from a Markov chain, let $\pi_i(X)$ be the exit sequence of s_i (the symbols following s_i) for $1 \leq i \leq m$. Assume that $[\Lambda_1, \Lambda_2, ..., \Lambda_n]$ is an arbitrary collection of exit sequences such that Λ_i and $\pi_i(X)$ are permutations and they have the same last element for all $1 \leq i \leq m$. Then there exists a sequence $X' = x'_1x'_2...x'_N \in \{s_1, s_2, ..., s_m\}^N$ such that $x'_1 = x_1$ and $\pi_i(X') = \Lambda_i$ for all $1 \leq i \leq m$. For this X' , we have $x'_N = x_N$.*

V. CONCLUSION

In this paper, we addressed the problem of generating random-bit streams from i.i.d. sources with unknown distributions. First, we considered the case of biased coins and derived a simple algorithm to generate random-bit streams. This algorithm achieves the information-theoretic upper bound on efficiency. We showed that this algorithm can be generalized to generate random-bit streams from an arbitrary m -sided die with $m > 2$, and its information efficiency is also

asymptotically optimal. Furthermore, we demonstrated that by applying the (generalized) random-stream algorithm, we can generate random-bit streams from an arbitrary Markov chain very efficiently.

REFERENCES

- [1] M. Blum, “Independent unbiased coin flips from a correlated biased source: A finite state Markov chain,” *Combinatorica*, vol. 6, pp. 97–108, 1986.
- [2] T. M. Cover, J. A. Thomas, *Elements of Information Theory*, Second Edition, Wiley, July 2006.
- [3] P. Elias, “The efficient construction of an unbiased random sequence,” *Ann. Math. Statist.*, vol. 43, pp. 865–870, 1972.
- [4] S. Pae and M. C. Loui, “Optimal random number generation from a biased coin,” in *Proc. ACM-SIAM Symp. Discrete Algorithms*, pp. 1079–1088, 2005.
- [5] Y. Peres, “Iterating von Neumann’s procedure for extracting random bits,” *Ann. Statist.*, vol. 20, pp. 590–597, 1992.
- [6] B. Y. Ryabko and E. Matchikina, “Fast and efficient construction of an unbiased random sequence,” *IEEE Trans. Inform. Theory*, vol. 46, pp. 1090–1093, 2000.
- [7] P. A. Samuelson, “Constructing an unbiased random sequence,” *J. Amer. Statist. Assoc.*, pp. 1526–1527, 1968.
- [8] J. von Neumann, “Various techniques used in connection with random digits,” *Appl. Math. Ser.*, Notes by G.E. Forstyle, Nat. Bur. Stand., vol. 12, pp. 36–38, 1951.
- [9] H. Zhou and J. Bruck, “Efficient generation of random bits from finite state Markov chains,” *IEEE Trans. Inform. Theory*, vol. 58, pp. 2490–2506, Apr. 2012.
- [10] H. Zhou and J. Bruck, “A universal scheme for transforming binary algorithms to generate random bits from loaded dice,” Technical Report, Department of Electrical Engineering, California Institute of Technology, 2012.